

Normalization of Testosterone Levels After Testosterone Replacement Therapy Is Associated With Decreased Incidence of Atrial Fibrillation

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Background—Atrial fibrillation (AF) is the most common cardiac dysrhythmia associated with significant morbidity and mortality. Several small studies have reported that low serum total testosterone (TT) levels were associated with a higher incidence of AF. In contrast, it is also reported that anabolic steroid use is associated with an increase in the risk of AF. To date, no study has explored the effect of testosterone normalization on new incidence of AF after testosterone replacement therapy (TRT) in patients with low testosterone.

Methods and Results—Using data from the Veterans Administrations Corporate Data Warehouse, we identified a national cohort of 76 639 veterans with low TT levels and divided them into 3 groups. Group 1 had TRT resulting in normalization of TT levels (normalized TRT), group 2 had TRT without normalization of TT levels (nonnormalized TRT), and group 3 did not receive TRT (no TRT). Propensity score-weighted stabilized inverse probability of treatment weighting Cox proportional hazard methods were used for analysis of the data from these groups to determine the association between post-TRT levels of TT and the incidence of AF. Group 1 (40 856 patients, median age 66 years) had significantly lower risk of AF than group 2 (23 939 patients, median age 65 years; hazard ratio 0.90, 95% CI 0.81–0.99, $P=0.0255$) and group 3 (11 853 patients, median age 67 years; hazard ratio 0.79, 95% CI 0.70–0.89, $P=0.0001$). There was no statistical difference between groups 2 and 3 (hazard ratio 0.89, 95% CI 0.78–1.0009, $P=0.0675$) in incidence of AF.

Conclusions—These novel results suggest that normalization of TT levels after TRT is associated with a significant decrease in the incidence of AF. (*J Am Heart Assoc.* 2017;6:e004880. DOI: 10.1161/JAHA.116.004880.)

Key Words: atrial fibrillation • testosterone • testosterone replacement therapy

Atrial fibrillation (AF) is the most common cardiac arrhythmia worldwide, with significant morbidity, mortality, and financial burden.^{1–3} In the United States, 2.7 to 6.1 million people had AF in 2010, and that number is expected to double by 2050.^{4,5} The prevalence and incidence

of AF is known to increase with age and is higher in men than in women.⁶ Although the underlying mechanisms of this sex difference are still unclear, 1 preclinical and several small clinical studies have suggested that testosterone deficiency may play a role in the development of AF.^{7–9} It was also reported that anabolic steroid use was associated with an increased risk of AF.^{10,11}

Serum levels of testosterone decline with age.¹² Hypogonadism is estimated to be present in 38% of men aged >45 years.¹³ However, the effect of testosterone replacement therapy (TRT) on cardiovascular health is a topic of ongoing debate.^{14,15} Our group and others have reported that normalization of testosterone levels after TRT is associated with a significant decrease in all-cause mortality, myocardial infarction, and stroke.¹⁶ In addition, we found that TRT did not increase the incidence of deep venous thrombosis or pulmonary embolism.¹⁷

To date, no studies have investigated the effect of testosterone-level normalization on incidence of new AF in

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men after TRT. Patients receiving TRT are older, and advanced age increases the risk of AF; therefore, novel information on the effect of testosterone levels and TRT on the incidence of AF would be clinically significant. In this study, we investigated the incidence of AF in hypogonadal men with documented low testosterone levels. We compared the incidence of AF among patients who did not receive any TRT, those who received TRT that resulted in normalization of total testosterone (TT), and those who received TRT but that did not result in normal TT levels. This allowed us to evaluate the impact of nontreatment as well as the impact of adequate replacement therapy on AF incidence in this specific population. We hypothesized that normalization of TT levels after TRT would be associated with a significant decrease in the incidence of AF.

Methods

Data Source

The US Veterans Health Administration (VHA) provides care to veterans at >1400 establishments across the United States. Clinical data from these establishments are archived in the Corporate Data Warehouse and made available for research through the Veterans Administration Informatics and Computing Infrastructure. Information regarding data quality can be found at the US Department of Veterans Affairs (VA) Information Resource Center (<http://www.virec.research.va.gov>). We conducted a retrospective cohort study of male veterans who received medical care through the VHA during the period December 1999 to May 2014. The institutional review board of the Kansas City VA Medical Center approved the study. Because this study used retrospective database research, the requirement for informed consent was waived.

Ascertainment of TT Level

We observed wide variability in the units used to express testosterone test results obtained from a large number of laboratories in the VA health system over the long follow-up period involved in this study. In addition, reported TT levels lacked uniform normal laboratory reference ranges. Thus, TT was considered low when the reported value was less than the lower limit of the normal laboratory reference range of the test result. We used this method rather than a discrete cutoff value because we found that reference ranges and reporting units varied with assays used at different facilities across the VA.^{18,19} Even in the same hospital, the assay used changed over time. A lack of standardization for testosterone levels and other tests using stoichiometric measurements further adds to differences in the

presentation of test results.^{20,21} Consequently, we classified each test result as low or normal TT based on the respective normal laboratory reference range reported with the test result, as this was the most accurate method to minimize the effect of use of multiple assays.

Exposure Variable

We determined the use of TRT from patient medication prescription records. Patients who received any form of TRT (injection, gel, or patch) were considered to as treated. Untreated patients did not receive any form of TRT. Those who received TRT were categorized as normalized treated or nonnormalized treated depending on whether their TT levels improved above the normal laboratory reference range while on treatment or remained low.

Outcome Variable

Primary outcome measure was the incidence of AF (International Classification of Diseases, Ninth Revision [ICD-9] code 427.31).

Confounding Factors

Confounding measures were patient demographics; comorbidity, such as diabetes mellitus, hypertension, chronic obstructive pulmonary disease, obstructive sleep apnea, congestive heart failure, peripheral vascular disease, coronary artery disease, hyperthyroidism, rheumatic valve disease, nonrheumatic valve disease, structural heart disease, and cardiomyopathy; baseline body mass index; low-density lipoprotein; and use of aspirin, β -blockers, and statins. The ICD-9 Clinical Modification diagnosis codes were used to capture the coexisting conditions.

Study Population

Figure 1 presents the patient-selection process.

Inclusion criteria

We included patients whose TT levels in the first test were lower than the respective normal laboratory reference range.

Exclusion criteria

We excluded (1) women, (2) patients who received TRT before the first results showing low TT, (3) those who had AF before the first day of this study, (4) those with normalized TT levels on second testing without evidence of TRT, and (5) those with history of pericarditis and myocarditis.

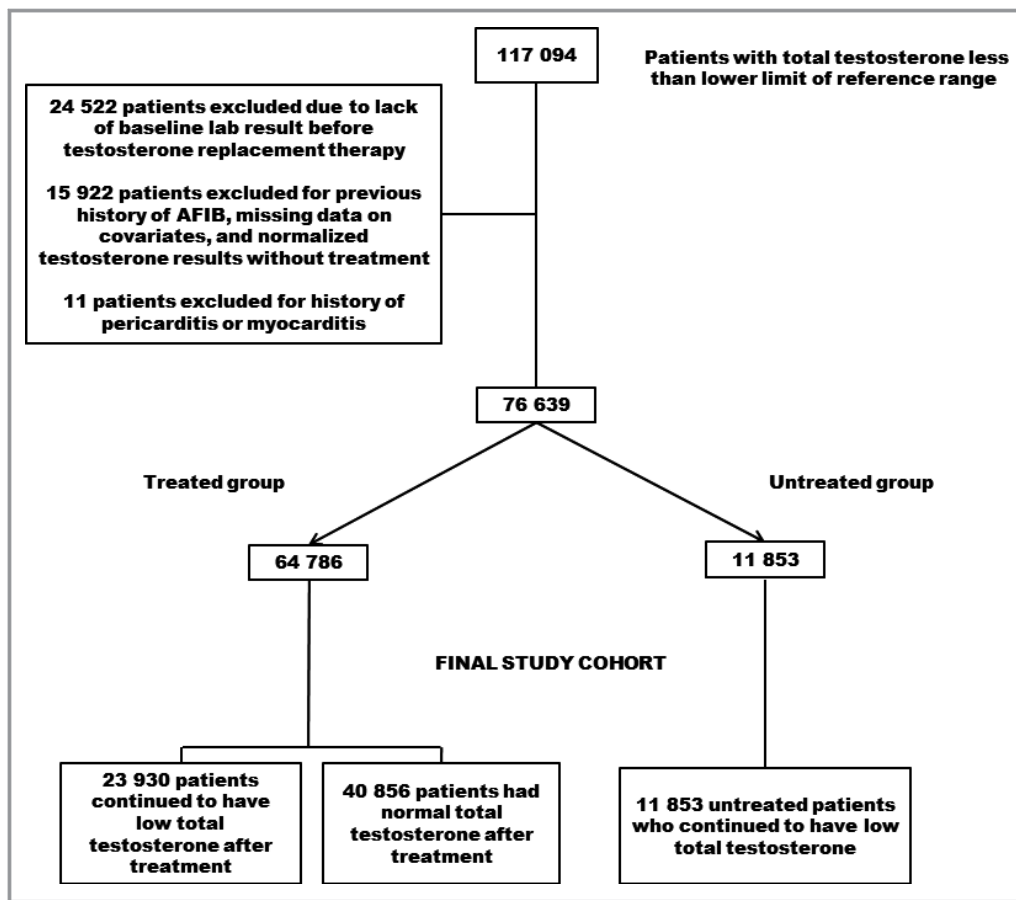


Figure 1. Methodology and patient-selection process, with inclusion and exclusion criteria to obtain the final cohort. AFib indicates atrial fibrillation.

Statistical Analysis

To enhance the robustness of our analysis and to ensure that the groups being compared were well matched, we utilized propensity score–weighted stabilized inverse probability of treatment weights (IPTW). Use of IPTW allowed us to keep the largest number of participants in the study after matching while using the propensity scores to achieve a balance between each pair of subgroups studied. Using this method, participants are weighted by the inverse probability of their treatment status. We preferred a stabilized IPTW approach to regular IPTW or propensity score matching because stabilized IPTW enhances inclusion of the largest number of patients in the analysis, whereas propensity score matching may result in a decreased sample size. In addition, stabilized IPTW significantly corrects for the instability in estimated treatment weights that potentially results from the use of regular IPTW for participants with very low probability of treatment.²² Furthermore, propensity scores were used to correct for potential systematic differences between treated and untreated participants. Each participant's propensity scores for receiving TRT were computed and adjusted for the covariates in a logistic regression

analysis. The covariates included were age, body mass index, hypertension, diabetes mellitus, chronic obstructive pulmonary disease, obstructive sleep apnea, congestive heart failure, peripheral vascular disease, coronary artery disease, hyperthyroidism, rheumatic valve disease, nonrheumatic valve disease, structural heart disease, cardiomyopathy, low-density lipoprotein, and use of aspirin, β -blockers, and statins.

We computed the incidence of AF in each subgroup. The χ^2 test and Student *t* test were used to compare normally distributed baseline characteristics of patients. Nonparametric tests were used for nonnormally distributed variables. Univariate and multivariable Cox proportional hazards regression analyses were conducted to assess the differences between groups. Continuous variables were reported as mean and standard deviation, categorical variables were reported as percentages. We also applied the stabilized IPTW to obtain Kaplan–Meier survival curves and to compare event-free survival time between groups, along with log-rank *P* value. Patients who did not develop AF episode were censored at the end date of the observation period. Patients who died were censored at the time of death. This enabled us to capture the

time contributed to the study by all enrolled patients. We assumed that censoring was random and unrelated to the likelihood of developing AF; therefore, patient failure times or time to incident AF would be the same as if these participants were actually observed for the study period. SAS Enterprise Guide 7.1 supported on SAS 9.4 (SAS Institute) was used for statistical analyses, with TRT as a time-varying exposure variable. The study hypothesis was tested at a 2-sided level of significance with $P < 0.05$.

Results

Cohort Description

We identified 117 094 patients with low TT. Of these, we excluded 24 522 patients who did not have baseline TT results before initial TRT. We also excluded 15 922 patients who had previous history of AF, those who had missing data on the baseline covariates, and those whose achieved normalized TT without any record of treatment, and 11

Table 1. Baseline Characteristics of All Patients in the Study Before and After Propensity Matching

	Baseline Characteristics of Cohorts			P Values for Unmatched and IPTW-Matched Cohorts					
	Normalized Treated (n=40 856)	Nonnormalized Treated (n=23 930)	Untreated (n=11 853)	₁ P Value*	₁ P Value [†]	₁ P Value [‡]	₂ P Value*	₂ P Value [†]	₂ P Value [‡]
Age ≥50 y, n (%)	36 590 (89.6)	21 260 (88.8)	10 687 (90.2)	0.0044	0.0567	0.0001	0.9910	0.9910	0.9958
Age, median (y)	66.0	65.0	67.0						
Body mass index ≥30	27 034 (66.2)	16 614 (69.4)	7583 (64.0)	<0.0001	<0.0001	<0.0001	0.9306	0.9306	0.9498
Body mass index, mean (SD)	33.1 (6.6)	33.7 (6.9)	32.9 (6.8)						
Follow-up time (y), mean (SD)	6.0 (3.1)	4.4 (2.9)	4.5 (2.9)						
Hypertension, n (%)	7252 (17.8)	4931 (20.6)	2108 (17.8)	<0.0001	0.9313	<0.0001	0.9406	0.9406	0.9403
Diabetes mellitus, n (%)	12 666 (31.0)	8787 (36.7)	3886 (32.8)	<0.0001	0.0002	<0.0001	0.9925	0.9925	0.9260
Chronic obstructive pulmonary disease, n (%)	489 (1.2)	427 (1.8)	188 (1.6)	<0.0001	0.0009	0.1744	0.9478	0.9478	0.9121
Obstructive sleep apnea, n (%)	761 (1.9)	671 (2.8)	259 (2.2)	<0.0001	0.0249	0.0005	0.9983	0.9983	0.8923
Congestive heart failure, n (%)	643 (1.6)	589 (2.5)	288 (2.4)	<0.0001	<0.0001	0.8557	0.9937	0.9937	0.8371
Peripheral vascular disease, n (%)	354 (0.9)	291 (1.2)	146 (1.2)	<0.0001	0.0003	0.8987	0.9851	0.9851	0.8857
Coronary artery disease	2258 (5.5)	1695 (7.1)	697 (5.9)	<0.0001	0.1406	<0.0001	0.9646	0.9646	0.9668
Depression, n (%)	3477 (8.5)	2187 (9.1)	780 (6.6)	0.0062	<0.0001	<0.0001	0.9594	0.9594	0.8799
Cardiomyopathy, n (%)	840 (2.1)	587 (2.5)	327 (2.8)	0.0009	<0.0001	0.0844	0.9903	0.9903	0.9613
Rheumatic valve disease, n (%)	484 (1.2)	380 (1.6)	181 (1.5)	<0.0001	0.0033	0.6624	0.9913	0.9913	0.8648
Nonrheumatic valve disease, n (%)	1427 (3.5)	938 (3.9)	577 (4.9)	0.0052	<0.0001	<0.0001	0.9814	0.9814	0.8922
Structural heart disease, n (%)	758 (1.9)	531 (2.2)	266 (2.2)	0.0014	0.0069	0.8792	0.9525	0.9525	0.9111
Bacterial endocarditis, n (%)	47 (0.1)	47 (0.1)	17 (0.1)	0.9433	0.4346	0.5069	0.9533	0.9533	0.9734
Hyperthyroidism, n (%)	559 (1.4)	273 (1.1)	180 (1.5)	0.0131	0.2202	0.0026	0.9451	0.9451	0.9662
LDL >100 mg/dL, n (%)	21 082 (51.6)	11 458 (47.9)	5778 (48.8)	<0.0001	<0.0001	0.1229	0.9476	0.9476	0.8929
Concomitant therapy									
Antiplatelet agents, n (%)	11 907 (29.1)	7535 (31.5)	3550 (30.0)	<0.0001	0.0895	0.0031	0.9708	0.9708	0.9679
β-Blockers, n (%)	15 336 (37.5)	9999 (41.8)	4579 (38.6)	<0.0001	0.0304	<0.0001	0.9765	0.9765	0.9549
Statins, n (%)	24 352 (59.6)	15 141 (63.2)	7134 (60.2)	<0.0001	0.2547	<0.0001	0.9619	0.9619	0.9592

₁P indicates unmatched P value; ₂P, matched P value; IPTW, inverse probability of treatment weights; LDL, low-density lipoprotein.

*Normalized treated vs nonnormalized treated.

[†]Normalized treated vs untreated.

[‡]Nonnormalized treated vs untreated.

patients were excluded for history of pericarditis and/or myocarditis. The final cohort for the study comprised 76 639 participants who were later divided into 40 856 normalized-treated and 23 930 nonnormalized-treated, and 11 853 untreated patients.

We excluded those with normalized TT without any record of treatment because we could not rule out the possibility of non-VA prescriptions as a reason for this finding. To prevent misclassification bias, these participants with spuriously normalized testosterone levels were excluded.

Baseline Characteristics of Study Participants

Table 1 presents the baseline variables for the unmatched and stabilized IPTW-matched cohorts. Median ages at enrollment were 66.0, 65.0, and 67.0 years for groups 1, 2, and 3, respectively. The mean body mass index (kg/m^2) at enrollment was 33.1 (SD 6.6), 33.7 (SD 6.9), and 32.9 (SD 6.8) for groups 1, 2, and 3, respectively. Mean follow-up time was 6.0 ± 3.1 years for the normalized TRT group, 4.4 ± 2.9 years for the nonnormalized TRT group, and 4.5 ± 2.9 years for the untreated group.

Relationship Between TRT and AF

The incidence of AF was as follows: group 1 (normalized TRT), 497 per 100 000 person years; group 2 (nonnormalized TRT), 626 per 100 000 person-years; group 3 (no TRT), 697 per 100 000 person-years. Table 2 presents the results of Cox proportional hazards regression analysis. Group 1 had significantly lower risk of AF than group 2 (hazard ratio 0.90, 95% CI 0.81–0.99, $P=0.0255$) and group 3 (hazard ratio 0.79, 95% CI 0.70–0.89, $P=0.0001$). There was no statistically significant decrease in AF risk between groups 2 and 3 (hazard ratio 0.89, 95% CI 0.78–1.009, $P=0.0675$).

The Kaplan–Meier curves (Figure 2A and 2B) showed that the normalized-TRT group was associated with significantly higher probability of AF-free survival compared with the no-TRT group (log-rank $P=0.0001$) (Figure 2A) or the nonnormalized-TRT group (log-rank $P=0.03$) (Figure 2B). There was no significant difference regarding AF-free survival probability between the nonnormalized- and no-TRT groups (Figure 2C).

Discussion

This study highlights that normalization of TT levels by TRT is associated with a significantly lower incidence of AF in men with low initial TT levels and without prior history of AF compared with a matched group of participants who did not receive TRT. In addition, the incidence of AF was higher in participants who failed to achieve normal TT levels after TRT (likely due to inadequate replacement dose or

Table 2. The Risk of Incident AF Was Significantly Lower in the Normalized Treated Group Compared With the Nonnormalized Treated and Untreated Groups

Hazard Ratios for Atrial Fibrillation			
	HR	95% CI	P Value
Normalized treated vs untreated (reference)			
Univariate n=40 856 vs n=11 853	0.713	0.635–0.801	<0.0001
Propensity matched (stabilized IPTW) n=40 857 vs n=11 853	0.792	0.702–0.893	0.0001
Normalized treated vs nonnormalized treated (reference)			
Univariate n=40 856 vs n=23 930	0.800	0.727–0.880	<0.0001
Propensity matched (stabilized IPTW) n=40 859 vs n=23 928	0.896	0.813–0.987	0.0255
Nonnormalized treated vs untreated (reference)			
Univariate n=23 930 vs n=11 853	0.895	0.788–1.017	0.0881
Propensity matched (stabilized IPTW) n=23 930 vs n=11 855	0.888	0.782–1.009	0.0675

The risk was also lower in non-normalized treated group compared to the untreated group but lacks statistical significance. AF indicates atrial fibrillation; HR, hazard ratio; IPTW, inverse probability of treatment weights.

noncompliance) compared with those with normalization of TT levels following TRT. Our retrospective study describes the largest cohort of such patients and the longest follow-up of TRT to date.

AF is the most common cardiac arrhythmia encountered in clinical practice.³ Significant sex- and age-based differences exist in the prevalence of AF. Men have a 1.5-fold higher risk of developing AF compared with women after adjustment for several confounding factors.²³ In addition, age-specific differences in the prevalence of AF were more pronounced in men compared with women.^{24,25} Increase in the incidence of AF in men appears to coincide with a decrease in serum testosterone levels related to aging.²⁶ In a preclinical study, Tsuneda et al used orchietomized male Sprague–Dawley rats to show that electrical stimulus caused a significant increase in repetitive atrial responses without changes in other electrophysiological properties. They also found that increased atrial responses were abolished by administration of testosterone. The authors proposed that deficiency of testosterone was associated with increased atrial arrhythmogenicity and AF.⁸ It is also important to note that Tsai et al showed that in aged rabbits, testosterone replacement increased pulmonary vein and left atrium arrhythmogenesis by enhancing adrenergic activity.²⁷

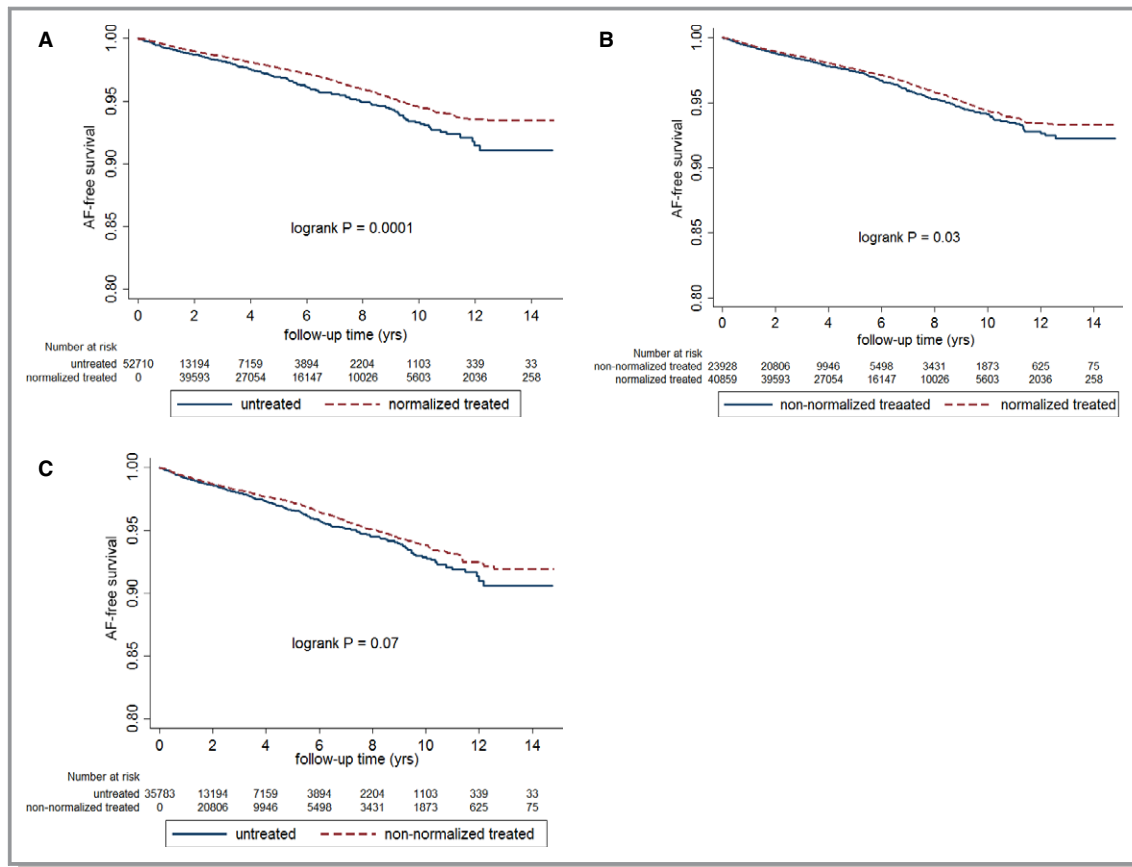


Figure 2. A, Kaplan–Meier curve showing atrial fibrillation–free (AF-free) survival probability between normalized treated and untreated participants. Normalized treated patients had higher AF-free survival probability than untreated participants. B, Kaplan–Meier curve showing AF-free survival probability between normalized and nonnormalized treated patients. Normalized treated patients had higher AF-free survival probability than the nonnormalized treated patients. C, Kaplan–Meier curve showing AF-free survival probability between nonnormalized treated and untreated participants. There was no statistically significant difference regarding AF-free survival probability between nonnormalized treated and no-TRT groups.

The mechanisms by which normal levels of testosterone might reduce risk of AF development remain to be established and are likely to be multifactorial. It has been proposed that calcium leakage caused by phosphorylation of ryanodine receptor type 2 from the sarcoplasmic reticulum plays a pivotal role in pathogenesis and progression of AF.²⁸ Tsuneda et al, in their orchietomized male rat model, showed that the difference in atrial response was associated with increased ryanodine receptor type 2 and sodium-calcium exchange, which were attenuated by administration of testosterone. This study suggests an antiarrhythmogenic property for testosterone.⁸ In recent years, inflammation has been implicated in the pathogenesis of AF.^{29–33} Inflammatory markers such tumor necrosis factor α , interleukin 6, interleukin 2, and C-reactive protein were all found to be elevated in patients with AF.^{34–40} Although data are not consistent, several authors have reported decreased levels of tumor necrosis factor α , interleukin 6, C-reactive protein, and other proinflammatory cytokines after TRT.^{41–45} Consequently, the anti-

inflammatory property of testosterone is a potential mechanism for the observed reduction in the risk of AF. However, additional basic science as well as clinical studies will be needed to appropriately define the mechanisms of testosterone-induced reduction in the risk of AF.

The first clinical study that suggested an association between low testosterone and AF was a single-center study of 58 male participants with lone AF. Participants were age and sex matched. The mean age of the participants was 46.1 ± 9.7 for lone AF and 45.2 ± 8.6 years for control. This study found that mean testosterone levels were significantly lower in participants with lone AF.⁷ Recently, a study included 1251 participants (aged 68.0 ± 8.2 years) and assessed the 10-year risk of AF by utilizing multivariable-adjusted hazard models.⁴⁶ Participants were stratified into ages 55 to 69 years ($n=786$), 70 to 79 years ($n=351$), and ≥ 80 years ($n=114$). The authors found that the risk of AF increased by 1.3- and 3.5-fold in men aged 55 to 69 years and ≥ 80 years, respectively, with every 1-SD decrease in TT

levels. They also observed a trend of increased AF in men aged 70 to 79 years with decreased testosterone, but that was statistically not significant (hazard ratio 1.14, 95% CI 0.91–1.44).

Compared with the studies discussed above, the present study included 76 639 participants with a median age of 65 to 66 years. Those with low initial TT were followed for 4 to 6 years, with or without TRT treatment. We found that the incidence of AF was significantly higher among participants with low TT levels who received no treatment compared with those who achieved normal TT levels after TRT. These results, in conjunction with previous studies, strengthen the hypothesis that testosterone deficiency may potentiate the risk of developing AF. Stroke is a major disabling and sometimes fatal complication of AF. Our previous study,¹⁶ which showed a significant decrease in the incidence of stroke with testosterone normalization after TRT, strengthens the current findings.

Although studies^{7,46} have demonstrated association between low testosterone and an increased risk of AF, additional studies will be needed to evaluate the strength of the association between hypogonadism and AF compared with the strength of other established risk factors. It is possible that low endogenous testosterone is associated with these risk factors and may not be a part of the causal pathway. The TOM (Testosterone in Older Men) trial, which was designed to determine the effects of TRT on lower extremity strength and physical function in older men with limitations in mobility and low serum levels of TT or free T, found 3 AF episodes in the testosterone group compared with none in the placebo group.¹⁵ In a recent study, Snyder et al reported 11 hospitalizations for arrhythmias in the testosterone group compared with only 7 in the placebo group.⁴⁷ In addition, 2 case reports suggest a risk of AF by anabolic steroid use or abuse.^{10,11}

It is worth mentioning that the US Food and Drug Administration recently approved changes to the labeling instructions on testosterone products that provides warning regarding the abuse potential of testosterone and the serious adverse outcomes, especially those related to cardiovascular and mental health, that have been reported in association with testosterone/anabolic androgenic steroid abuse.⁴⁸

Study Limitations

As an observational study, certain limitations are inherent to the design. Unmeasured confounding or hidden biases might be present. Another limitation of our study is the lack of randomization. Our database does not permit us to ascertain the time of the day when TT levels were drawn. Although blood samples are normally collected during morning hours in the VA health care system, some patients may have had their

samples drawn later in the day, resulting in potential underestimation of TT levels. Another limitation of observational studies is the validity of inclusion and exclusion parameters. However, these criteria were determined using ICD-9 codes, and the VA cohort ICD-9 codes have been validated for determining outcomes.¹⁴ Despite this, we recognize that in some patients in our study, AF episodes may have been missed either because patients did not recognize having an AF episode or an AF episode spontaneously stopped before a patient sought medical help.

Conclusion

To the best of our knowledge, this study is the first and the largest to evaluate the association between TRT and the incidence of AF. Results from our study show that low TT levels are associated with increased incidence of AF. The largest and significant decrease in the incidence of AF was observed in the group with normalized TT levels following TRT. Future prospective randomized controlled trials are needed to validate these results.

Disclosures

None. The contents of this article are those of authors and do not necessarily reflect the position and policy of the Department of Veterans Affairs or the United States Government.

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